

Bayesian Networks

4. Exact Inference / Variable Elimination

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Bayesian Networks



1. Inference in Probabilistic Networks

2. Variable elimination

studfarm example



Figure 1: Genealogical structure for the horses in the studfarm example [Jen01, p. 47].

	aa	aA	AA
aa	(1, 0, 0)	(0.5, 0.5, 0)	(0, 1, 0)
aА	(0.5, 0.5, 0)	(0.25, 0.5, 0.25)	(0, 0.5, 0.5)
AA	(0, 1, 0)	(0, 0.5, 0.5)	(0, 0, 1)

Figure 2: p(Child | Father, Mother) for genetic inheritance. The numbers are the probabilities for (aa, aA, AA) [Jen01, p. 47].

Variable *disease* with three states:

pure (aa) carrier (aA) sick (AA)

Genalogic graph becomes bayesian network if

(i) each non-root vertex has conditional probability distribution

p(child|father, mother)

as given in fig. 2,

(ii) each root vertex has probability distribution

$$p(aa) = .99, p(aA) = .01, p(AA) = .0$$

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Bayesian Networks / 1. Inference in Probabilistic Networks

studfarm example





	aa	aA	AA
aa	(1, 0, 0)	(0.5, 0.5, 0)	(0, 1, 0)
аA	(0.5, 0.5, 0)	(0.25, 0.5, 0.25)	(0, 0.5, 0.5)
AA	(0, 1, 0)	(0, 0.5, 0.5)	(0, 0, 1)

GwenFigure 4: p(Child | Father, Mother) for genetic inheritance. The numbers are the probabilities for (aa, aA, AA) [Jen01, p. 47].

Figure 3: Genealogical structure for the horses in the studfarm example [Jen01, p. 47].

father	aa			aA			AA]	father	aa		aA]					
mother	aa	aΔ	ΔΔ	aa	aΔ	ΔΔ	aa	aΔ	ΔΔ		mother	aa	aΔ	aa	aΔ		father	aa		aA	
mounor	uu	u/ (703	uu	un	703	uu	u/ (703		mounor	uu	un	uu	u/		mother	22	aΔ	22	aΔ
22	1	5	Δ	5	25	Δ		Δ	Δ		22	1	5	5	25		mound	uu	an	uu	an
aa	'	.0	0		.20	0		0	0		aa	1	.5		.20		22	1	5	5	1
<u>م</u> ۸		Б	-1	5	Б	Б	1	Б	Δ		<u>م</u> ۸		Б	5	Б		aa		.5		3
aA		.0	1	.5	.0	.0		.0	0		aA		.5	.0	.0		~^		5	5	2
	0	0	0	0	25	5	0	5	1			0	0	0	25		aA	0	.5	.o	$\overline{3}$
7.01		0	0	0	.20	.0		.0			7.01	0	0		.20	1					

Figure 5: p(child | father, mother) in general (left), if father and mother cannot be sick (middle), and if child cannot be sick either (right).

studfarm example / "forward inference"



$$p(aa) = 0.99 \cdot 0.99$$

 $p(aA) = +2 \cdot \frac{1}{2} \cdot$

 $+\frac{1}{2}$

=0.00995

=0.000025

 $p(AA) = +\frac{1}{4}.$

$$+2\cdot\frac{1}{2}$$
 $0.99\cdot0.01$

 $0.99 \cdot 0.01$

 $0.01 \cdot 0.01$

 $0.01 \cdot 0.01$

$$+\frac{1}{4}$$
. $0.01 \cdot 0.01$
=0.990025

Figure 6: Genealogical structure for the horses in the studfarm example [Jen01, p. 47].

father	aa		aA	
mother	aa	аA	aa	аA
aa	1	.5	.5	.25
aA	0	.5	.5	.5
AA	0	0	0	.25

Figure 7: p(child | father, mother) if father and mother cannot be sick.

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studfarm example / "forward inference"

Brian Ann Cecily Cai 1.001 Cai 1.00 Cai Cai 1.00 1.001 Cai 1.001 Pur 99.001 Pur 99.00 Pur 99.00 Pur 99.00 Pur 99.00 Dorothy Gwenn Cai Cai Cai 1.00 1.001 1.00 Cai 1.001 Pur 99.001 Pur 99.00 Pur 99.00 Pur 99.001 Henry Irer Cai 0.931 Cai 1.00 Pur 99.07 Pur 99.001 Johr 0.04 Cai 0.881

Figure 8: Probabilities without evidence. [Jen01, p. 49]

Pur 99.08







studfarm example / "backward inference"



Figure 9: Genealogical structure for the horses in the studfarm example [Jen01, p. 47].

father	aa		aA	
mother	aa	аA	aa	аA
aa	1	.5	.5	.25
aA	0	.5	.5	.5
AA	0	0	0	.25

Figure 10: p(child | father, mother) if father and mother cannot be sick.

(iii) Henry and Irene are carrier (aA) with p = 1.

If only Fred, Dorothy, Erik, and Gwen existed, we could further infer that for each of them

$$p(aa) = \frac{1}{3}, \quad p(aA) = \frac{2}{3}$$

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studfarm example / "backward inference"



Figure 11: Probabilities given evidence that John is sick (AA). [Jen01, p. 49]

Evidence

Definition 1. Let *V* be a set of variables. The set

$$\mathcal{E} := \{ E \subseteq \bigcup_{v \in V} \{v\} \times \operatorname{dom}(v) \, | \, \forall (v, c), (v, c') \in E : c = c' \}$$

is called **space of evidence of** V. Evidence is a setting of variables to spe-An element $E \in \mathcal{E}$ is called **evidence of** cific values. "Fuzzy" or "uncertain evidence" that assigns probabilities to the V. We call $dom(E) := \{v \in V \mid \exists c \in dom(v) : (v, c) \in E \text{ plifferent values of the variables, is not } v \in V \mid \exists c \in dom(v) : (v, c) \in E \text{ plifferent values of the variables, is not } v \in V \mid \exists c \in dom(v) : (v, c) \in E \text{ plifferent values of the variables, } v \in V \mid \exists c \in dom(v) : (v, c) \in E \text{ plifferent values of the variables, } v \in V \mid \exists c \in dom(v) : (v, c) \in E \text{ plifferent values of the variables, } v \in V \mid \exists c \in dom(v) : (v, c) \in E \text{ plifferent values of the variables, } v \in V \mid dv \in V | v \in V | v \in V |$ handled here. the set of evidential variables and for each evidential variable $v \in dom(E)$ we call the unique $E_v := c \in \operatorname{dom}(v)$ with $(v, c) \in E$ its (evidential) value. Evidence E corresponds to the probability distribution epd_E : $\operatorname{dom}(v) \to \mathbb{R}^+_0$ $v \in \operatorname{dom}(E)$

 $(x)_{v \in \operatorname{dom}(E)} \quad \mapsto \begin{cases} 1, & \text{if } \forall v : (v, x) \in E \\ 0, & \text{otherwise} \end{cases}$

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Evidence / example

Example 1. Let $V := \{A, B, C, D\}$ and $dom(A) := dom(B) := \{0, 1\},$ $dom(C) := \{0, 1, 2\}$ and $dom(A) := \{0, 1, 2, 3\}.$

Then

$$E := \{ (A, 1), (C, 2) \}$$

is an evidence with the evidential variables A and C. The evidential variable A has value 1, the variable C value 2.

The probability distribution corresponding to ${\cal E}$ is

$$\operatorname{epd}_E(A=1, C=2) = 1$$

and

$$\operatorname{epd}_E(A=a,C=c)=0$$

for all other values a of A or c of C.





Entering evidence



Let V be a set of variables and q be a potential on a subset of V. Let E be evidence of V.

We call

$$q_E : \prod_{\substack{v \in \operatorname{dom}(q) \setminus \operatorname{dom}(E) \\ (x)_{v \in \operatorname{dom}(q) \setminus \operatorname{dom}(E)}} \operatorname{dom}(v) \to \mathbb{R}_0^+ \to q(x, E)$$

with

$$(x, E)(v) := \begin{cases} x_v, & \text{if } v \in \operatorname{dom}(q) \setminus \operatorname{dom}(E) \\ E_v, & \text{if } v \in \operatorname{dom}(E) \end{cases}$$

the potential q given evidence E.

If q is a JPD, then q_E is the probability distribution on the non-evidential variables $dom(q) \setminus dom(E)$ for outcomes that conform to E (i.e., have value E_v for each variable $v \in dom(E)$).

Warning: q_E should not be confused with the conditional probability distribution $q^{|\operatorname{dom}(E)}$. In sloppy notation for E = $\{(v_1, c_1), \ldots, (v_n, c_n)\}$:

$$q_E = q(x, v_1 = c_1, \dots, v_n = c_n)$$

and

$$q^{|\operatorname{dom}(E)|} = q(x | v_1, \dots, v_n)$$

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Inferencing



Given a JPD p on a set of variables V and evidence E on V.

We distinguish three types of inference targets:

(i) a single variable: For a given variable $v \in V$ infering v based on E w.r.t. p means to compute

$$p(v|E) = \frac{p(v,E)}{p(E)} \sim p(v,E)$$

or (more exactly) $(p_E)^{\downarrow v \mid \emptyset}$.

(ii) several variables separately: For a given set of variables W ⊆ V infering W separately based on E w.r.t. p means to compute

$$\begin{split} p(v|E) &= \frac{p(v,E)}{p(E)} \sim p(v,E), \quad \forall v \in W \\ \text{or } (p_E)^{\downarrow v | \emptyset} \end{split}$$

(iii) joint distribution of several variable For a given set of variables $W \subset V$

infering the marginal W based on E w.r.t. p means to compute

$$p(W|E) = \frac{p(W, E)}{p(E)} \sim p(W, E)$$

or
$$(p_E)^{\downarrow W \mid \emptyset}$$

Normalizing is necessary, as p_E in general is not a probability distribution, even if p is.

Bayesian Networks / 1. Inference in Probabilistic Networks

Inferencing / JPD as one large table

If p is given as one large table, infering the marginal W based on E means

- (i) select the subtable indexed by E,
- (ii) aggregate to W, i.e., sum over all variables $V \setminus \text{dom}(E) \setminus W$,

If we observe the evidence V = Y, then

(iii) normalize.

Pain	Y				Ν			
Weightloss	Y		Ν		Y		N	
Vomiting	Y	Ν	Y	Ν	Y	Ν	Y	Ν
Adeno Y	220	220	25	25	95	95	10	10
Ν	4	9	5	12	31	76	50	113

Figure 12: JPD p given as one large table.

Pain	Y		N	
Weightloss	Y	Ν	Y	Ν
Adeno Y	220	25	95	10
Ν	4	5	31	50

Figure 13: Subtable for $E = \{(V, Y)\}$: distribution p_E before normalization.

$$p(\text{adeno}=Y|V=Y) = \sum_{w,q} p(\text{adeno}=Y, W=w, P=q|V=Y)$$
$$= \frac{220 + 25 + 95 + 10}{224 + 30 + 126 + 60} = \frac{350}{440} = 0.80$$

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Bayesian Networks / 1. Inference in Probabilistic Networks



Inferencing / JPD as product of potentials

If p is given as product of potentials, i.e.,

$$p := (\prod_{q \in Q} q)^{|\emptyset|}$$

the problem becomes more interesting.

Naive approach: we reduce the problem to inference w.r.t. p as one large table by explicitly computing p and then doing inference as on the former slide, actually computing

$$(p_E)^{\downarrow W|\emptyset} = (((\prod_{q \in Q} q)^{|\emptyset})_E)^{\downarrow W|\emptyset}$$

Naive approach $_2$: we

- (i) enter evidence in the factors first, i.e., compute q_E , and then
- (ii) compute p_E as product of the q_E 's

$$(p_E)^{\downarrow W|\emptyset} = ((\prod_{q \in Q} q_E)^{\downarrow W|\emptyset})$$

product of potentials / naive approach



Pain Y	.52	Weit	hloss `	Y .75	Vc	miting	Y .4	4
	P)		V)	v)		V		
Pain	Y		(-	•/	N			
Weightloss	Y		N		Y		N	
Vomiting	Y	Ν	Y	Ν	Y	Ν	Y	Ν
Adeno Y	.982	.961	.833	.676	.754	.556	.167	.081

Figure 14: Bayesian Network for adeno JPD.

Pain	Y				N			
Weightloss	Y		N		Y		Ν	
Vomiting	Y	Ν	Y	Ν	Y	Ν	Y	N
Adeno Y	.169	.210	.048	.049	.119	.112	.009	.005
Ν	.003	.009	.010	.024	.039	.090	.044	.062

Figure 15: JPD of Bayesian Network for adeno JPD.

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Bayesian Networks / 1. Inference in Probabilistic Networks

product of potentials / naive approach

Pain	Y				N			
Weightloss	Y		N		Y		N	
Vomiting	Y	Ν	Y	Ν	Y	Ν	Y	Ν
Adeno Y	.169	.210	.048	.049	.119	.112	.009	.005
N	.003	.009	.010	.024	.039	.090	.044	.062

Figure 16: JPD p given as one large table.

Pain	Y		N	
Weightloss	Y	Ν	Y	Ν
Adeno Y	.169	.048	.119	.009
Ν	.003	.010	.039	.044

Adeno Y .345

Figure 17: Subtable for $E = \{(V, Y)\}$: distribution p_E before normalization.

	Ν	.096
Figure 18: Aggregate subtable for $E = \{(V, Y)\}$.		

If we observe the evidence V = Y, then

$$p(\text{adeno}=Y|V=Y) = \sum_{w,q} p(\text{adeno}=Y, W=w, P=q|V=Y)$$
$$= \frac{.345}{.345 + .096} = 0.782$$





Vounities 2003





Figure 19: Bayesian Network for adeno JPD.

Pain	Y				N			
Weightloss	Y		N		Y		N	
Vomiting	Y	Ν	Y	Ν	Y	Ν	Y	Ν
Adeno Y	.384	0	.109	0	.270	0	.020	0
Ν	.007	0	.023	0	.089	0	.100	0

Figure 20: JPD of Bayesian Network for adeno JPD with evidence V = Y entered.

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Bayesian Networks / 1. Inference in Probabilistic Networks

Overview of inference methods [Guo and Hsu 2001]

- (i) exact inference:
 - (a) Polytree algorithm
 - (b) conditioning
 - (c) clustering
 - (d) arc reversal
 - (e) variable elimination

- (ii) approximate inference:
 - (a) stochastic sampling
 - (b) model simplification
 - (c) search-based
 - (d) loopy propagation

(iii) symbolic inference.

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1. Inference in Probabilistic Networks

2. Variable elimination

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Bayesian Networks / 2. Variable elimination

Aggregating products

Doing inference using the naive $approach_2$,

$$(p_E)^{\downarrow W|\emptyset} = ((\prod_{q \in Q} q_E)^{\downarrow W|\emptyset})$$

we compute a large table as product of q_E and then aggregate to W.

Question: can we aggregate the factors and then multiply the aggregates?

$$(pq)^{\downarrow W} \stackrel{?}{=} p^{\downarrow W} q^{\downarrow W}$$

In general, this equation does not hold, as

$$(pq)^{\downarrow W}(x) = \sum_{y \in \prod_{X \in \operatorname{dom}(pq) \setminus W} \operatorname{dom}(X)} p(x, y)q(x, y)$$

but

$$(p^{\downarrow W}q^{\downarrow W})(x) = (\sum_{y \in \prod_{X \in \operatorname{dom}(p) \setminus W} \operatorname{dom}(X)} p(x, y)) \cdot (\sum_{y \in \prod_{X \in \operatorname{dom}(q) \setminus W} \operatorname{dom}(X)} q(x, y))$$

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naive But it is true for $dom(p) \cap dom(q) \subseteq W$, i.e., if p and q have no common variables except those in W.

> **Lemma 1.** Let p and q be two potentials on a subset of variables V. Let $W \subseteq V$ a subset of the variables. If $dom(p) \cap dom(q) \subseteq W$ then

$$(pq)^{\downarrow W} = p^{\downarrow W} q^{\downarrow W}$$

Variable elimination



We can make use of this observation for simplifying $(\prod_{q \in Q} q)^{\downarrow W}$:

(i) choose a variable $v \in V \setminus W$, clearly

$$(\prod_{q \in Q} q)^{\downarrow W} = ((\prod_{q \in Q} q)^{\downarrow cv})^{\downarrow W}$$

i.e., we can eliminate variable v first,

(ii) let

$$R := \{q \in Q \mid v \in \operatorname{dom}(q)\}$$

be all potentials which's domain contains v and

$$q' := \prod_{q \in R} q, \quad q_{\mathsf{rest}} = \prod_{q \in Q \setminus R} q$$

(iii) Then

$$\operatorname{dom}(q') \cap \operatorname{dom}(q_{\mathsf{rest}}) \subseteq V \setminus \{v\}$$

and thus

$$(\prod_{q \in Q} q)^{\downarrow W} = (q_{\mathsf{rest}} \cdot q'^{\downarrow cv})^{\downarrow W}$$

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Bayesian Networks / 2. Variable elimination

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Variable elimination

i inference-varelim(Q : set of potentials, W : set of variables) :² <u>while</u> $\bigcup_{q \in Q} \operatorname{dom}(q) \setminus W \neq \emptyset$ <u>do</u> choose $v \in \bigcup_{q \in Q} \operatorname{dom}(q) \setminus W$ arbitrarily Q := eliminate-variable(Q, v)4 5 <u>od</u> 6 <u>return</u> $(\prod_{q \in Q} q)^{|\emptyset|}$ 7 eliminate-variable(Q : set of potentials, v : variable) : $s \ R := \{q \in Q \mid v \in \operatorname{dom}(q)\}$ 9 $q' := (\prod_{q \in R} q)^{\downarrow cv}$ 10 return $Q \setminus R \cup \{q'\}$

Also known as **bucket elimination**.

Useful if the set W of variables to infer separately is small.



i.e., we replace the potentials R by

 $q'^{\downarrow cv}$

After this replacement, the variable v is eliminated from the potentials $Q' := Q \setminus R \cup \{q'^{\downarrow cv}\}.$

example







The conditional probabilities are

 $Q := \{ p(A), p(B|A), p(C|A), p(D|B),$ $p(E|B,C), p(F|C)\}$

We want to compute the marginal p(D)given evidence on F. Thus we add epd(F) to Q.

F, E, C, A, B

the following steps have to be performed:



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Bayesian Networks / 2. Variable elimination

example

For the elimination sequence

A, B, C, E, F

the following steps have to be performed:

$$p(A), p(B|A), p(C|A) \quad \{ \underbrace{A, B, C} \}$$

$$p(D|B), p(E|B, C) \quad \{ \underbrace{B, C, D, E} \} \quad \{ \underbrace{C, D, E, F} \} \quad p(F|C)$$

$$\{ \underbrace{D, \underbrace{E, F} \} \}$$

$$(\underbrace{D, \underbrace{F} \}) \quad epd(F)$$



References



[Jen01] Finn V. Jensen. Bayesian networks and decision graphs. Springer, New York, 2001.

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